

Export Orientation, Foreign Production, and the Growth of Medium Sized and Large Manufacturing Firms

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Abstract

Using data on large and medium sized Austrian manufacturing firms this paper analyzes empirically in which way a high degree of export orientation and foreign production abroad, as two modes of serving foreign markets, are related to the growth performance of domestic production. Robust median regressions do not suggest that foreign production adversely affects firm growth at the domestic location. Exports to non-EU countries, in combination with foreign production, especially contribute to growth at the domestic location.

Keywords: Growth of firms, multinational firms, exports
Jel.: L11, F23

1 Introduction

The increasing importance of foreign production by multinational firms (MNEs) must be seen as part of a firm's growth process. Especially, when exports are impeded by rising marginal costs of production or transportation the latter interpreted in a wide sense as the costs of lacking market proximity, tariff and non-tariff barriers, etc. (see Scherer et al., 1975; Buckley - Casson, 1981; Pfaffermayr, 1997), firms tend to gradually shift part of their production to foreign countries over the course of their expansion.

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The positive correlation between the level of foreign production and exports (or home production) found in many studies suggests a static, complementary relationship between foreign investment (production) and exports (see Markusen, 1995, for an overview). Simply referring to a static relationship, however, does not bring out the whole story: Although positively correlated in levels with domestic activities, depending on the form of foreign direct investment (FDI), production abroad may exert an important - maybe negative - impact on the growth potential of MNEs at their domestic location. In a small panel of MNEs in Austrian manufacturing, for example, the median firm reduced domestic employment by 3.1 percent (on average p.a.) between 1989 and 1997, but increased employment in foreign affiliates by 9.8 percent on average p.a.¹. A priori, this does not necessarily imply that domestic employment has been substituted by foreign affiliate employment. To derive valid conclusions on the relationship between domestic (employment) growth and that of foreign affiliates, one needs information on the counterfactual. The relationship could be defined as substitutional (complementary), if domestic employment c.p. were to have shrunk less (faster) or were to have grown faster (slower) in the absence of foreign production. One way to attain this information is to estimate an econometric model and test for the marginal impact of foreign production. To provide an estimate of such a counterfactual case is the main aim of the paper.

Summarizing the results, robust median regressions do not provide evidence that the large and medium sized Austrian manufacturing firms engaged in both exporting and running foreign affiliates, on average grow slower (or shrink faster) at their domestic location. In contrast, the econometric estimates suggest a positive effect and thus no significant substitution of domestic production from a dynamic point of view. Additionally, there is some evidence that the growth/size relationship is weaker for firms with foreign production and that the growth potential at the domestic location is probably higher if firms pursue a strategy based on both high export and foreign production (especially when exports to countries outside of Europe are concerned). However, the estimation results once more confirm the random nature of firm growth which makes it difficult to derive precise estimates, especially without detailed information on the form of the firms' FDI.

The paper proceeds as follows: the next section specifies the main hypothesis, discusses the econometric specification which will be estimated, describes the data and shortly refers to the robust estimation method implemented. Section 3 presents the estimation results. Section 4 concludes and summarizes the main findings.

¹ See Dell'mour, Egger, Gugler, Pfaffermayr and Wolfmayr-Schnitzer (2000) for more details.

2 The econometric specification, the data base and the estimation method

Econometric specification: Theoretical and empirical studies examining the determinants of the growth of firms (see, Geroski, 1998 for an overview) are mainly based on the observation that in many markets the distribution of firm size is extremely skewed. According to Gibrat's law of proportionate growth, firm size (measured in sales, employment or assets) grows according to a random walk. Geroski (1998) summarizes the empirical evidence by observing that "corporate growth is really very random". The theoretical work of Jovanovic (1992), Mitchell (2000) and others additionally suggests the importance of learning effects in the early phase of new start-ups, implying that the age of firms is inversely related to growth (see e.g. Evans, 1987 for empirical evidence).

Following the large body of empirical literature which tests Gibrat's law of proportionate firm growth, this paper uses a cross-section of firms and regresses the average log differences in firm size, here measured as employment at the domestic location, on initial size, export propensity, an indicator of foreign production and additional control variables. The age of firms is not available for all firms and could not be included.

The effect of foreign activities on domestic firm growth depends on the nature of FDI. Horizontal FDI saves transportation costs or compensates for lacking market proximity at the expense of additional fixed plant set-up costs. Generally, this form of FDI is based on a proximity/plant size trade-off (see Brainard, 1983, Markusen - Venables, 1998) and substitutes for exports. Therefore, negative effects on firm growth at the domestic location (here in terms of employment) have to be expected from this source. Vertical direct investments in services and distribution, in contrast, generate intra-firm exports and foster domestic firm growth. The same holds true if FDI induces additional headquarter services or aims at the fragmentation of the production process (Helpman - Krugman 1995). In the latter case, MNEs delocate some parts of the value added chain to foreign affiliates to exploit comparative advantages and specialize in tasks more suited to the domestic location. Here, firms are predicted to grow also at their domestic location, if fragmentation induces an increase in productivity and enables firms to gain market shares.

In this paper, the nature of foreign production remains unobserved. Instead, the existence of foreign production is measured by a time invariant dummy F , taking the value 1 if a firm operates a foreign affiliate. Therefore, a reduced form aimed at identifying the average effect of foreign production on domestic firm growth is estimated. Following Wagner - Bernard (1997) and Lui et. al. (1999)² among others, the growth equations also include time invariant indicators of export orientation (split-up into exports to the EU and exports to - mainly Eastern European - countries outside the EU; EXEU, EXNEU). In the present context, the interest lies in the impact of exports vs. foreign pro-

²Note, both studies did not find significant ex-post performance differences between exporters and non-exporters.

duction, as two modes of serving foreign markets, on growth performance at the domestic location. In particular, the common influence of both high export orientation and foreign production is tested by introducing interaction terms. Since especially fast growing firms - e.g. if they are in the possession of firm-specific assets - are likely to become direct investors over time, we look at the ex- post growth performance in the period following the measurement of foreign production and the propensity to export (i.e. at the beginning of the estimation period), in order to avoid endogeneity problems.

The inclusion of additional control variables is motivated by the large literature on firm growth, although data-availability limits possible choices. The average investment to sales ratio controls for investment propensity which is hypothesized as one of the factors accounting for producer heterogeneity (INQ, Doms et al., 1995; Lui et al., 1999). In order to control for the characteristics and its changes in the firms' market environment, such as industry specific growth patterns of demand (Geroski et al. 1997), a proxy for market growth³ (GSAGG) defined below, as well as regional dummies ('Bundesländer'), are additionally included. The latter account for differences in regional policies and also for varying forms of access to the foreign markets. For example, firms located near the borders to Eastern Europe may develop differently compared to those situated in the western part of Austria. Lastly, since foreign production is unlikely to be a substitute of production for the domestic market, growth in domestic sales additionally enters the growth equation as a control variable (GSHOME).

Using employment (B) as a measure of firm size, the econometric specification is formally based on:

$$\ln B = \beta_0 + \beta_1 \ln B_0 + \beta_2 \ln U_H + Z \quad (1)$$

where U_H denotes sales in the domestic market and Z captures labor demand induced (or substituted for) by the two different modes of serving foreign markets and the other controls. Since the only available information on foreign involvement is export propensity and an indicator on the existence of foreign production, the following reduced form is postulated for Z :

$$Z = f(F; EXEU; EXNEU; INQ; GSAGG) \quad (2)$$

Combining (1) and (2) gives the empirical specification which is estimated below for a cross-section of large and medium sized Austrian manufacturing firms:

$$\frac{\ln B_{T_i} - \ln B_{T_0}}{T_i - T_0} = \beta_0 + \beta_1 \ln B_{T_0;i} + \beta_2 GSHOME_i + \beta_3 F_{T_0;i} \ln B_{T_0;i} + \beta_4 F_{T_0;i} + \beta_5 EXEU_{T_0;i} + \beta_6 EXNEU_{T_0;i} + \beta_7 EXEU_{T_0;i} F_{T_0;i} + \beta_8 EXNEU_{T_0;i} F_{T_0;i} + \beta_9 INQ_i + \beta_{10} GSAGG_j + R_k + \epsilon_i \quad (3)$$

³ Industry dummies are not introduced, in order to avoid multicollinearity. If they are included instead of GSAGG; the estimation results do not change much. However, the standard errors of the estimated coefficients tend to be higher.

Subscript i refers to the firm, j indexes 18 two-digit industries, to which a firm can belong and k their region (9 'Bundesländer'). $GSHOME_i$ measures average growth of sales in the domestic market and is defined as $\frac{\ln U_{H,T_i} - \ln U_{H,T_0}}{T_i - T_0}$ with $U_H = U(1 - EXEU - EXNEU)$. $GSAGG_j$ and INQ_i are defined below. Since the underlying panel is not balanced, averages for each firm are taken over 2-4 observations.

Data: The empirical investigation uses the Investment Surveys⁴ covering the period 1996-1999, which WIFO conducts annually in co-operation with the EU DG-ECFIN (see European Commission, 1997, for details). The survey is not compulsory⁵ and does not comprise a random sample of enterprises, but rather follows the development of a fixed 'test group' over the course of years. Due to the relatively small number of newly founded firms in the survey, the unbalanced panel is not representative of the whole population; instead it mainly covers mature, medium sized and large firms. In addition, it has not always been possible to obtain a clear picture of mergers and acquisitions, i.e. growth rates may reflect both internal and acquisition growth.

The survey provides information on employment (B), sales from domestic production (U), and investments (I), all for the domestic location. Exports into the EU ($EXEU$) and into countries outside the EU ($EXNEU$) are both measured as shares of sales from domestic production. They are evaluated according to 11 categories and are valued at the middle of the corresponding interval. This wording of the questions greatly increases the questionnaire's general acceptance by the firms, with the disadvantage that some information is lost. The foreign production dummy (F) takes the value of 1; if the firm operates a foreign affiliate. Information on the size classes of amount of foreign production is not used. The investment sales ratio is defined as $INQ = \frac{1}{T_i - T_0} \ln(I=U)$. To control for overall demand growth, one ideally wants to include growth in apparent consumption in the foreign markets. Since this is not available, $GSAGG$ is calculated as the median of the average growth rate of the firm's sales over the period 1996-1999, in each two-digit industry. Note that this measure is based on all firms in the sample and not only on those used for the regressions.

The sample is restricted according to several criteria. First, only those firms can be analyzed which have provided information on all variables. Secondly, firms which do not decide between exporting and foreign production are excluded, because they either do not export or are relatively small (< 100 employees). Additionally, a few severe outliers with implausibly high or low growth rates for sales and/or employment (average log differences of employment above 0.5 or below -0.5), which most likely were the result of data-errors or due to

⁴ Although the WIFO Investment Survey basically defines the plant (Betrieb) as its unit of measurement, most larger enterprises prefer to respond at the corporate level. All observations referring to the plant level, which cannot decide to setup an affiliate abroad have been omitted. The survey asks for both realized values, which lag two years behind the survey date, as well as planned values. In the present study, the realized values are used; only the data for 1999 refer to planned values (4th plan, see European Commission, 1997).

⁵ So some firms dropped out of the sample not only because of an exit or takeover, but also because they no longer wished to participate.

significant mergers and acquisitions or restructuring, have been skipped. This leaves us with a cross-section of 351 firms from a total of 1144 firms.

Table 1

Table 1 provides definitions and summary statistics. The median of average employment growth per annum is 0.26 percent, and there is considerable variation, as illustrated by the growth rates for the first and third quartiles. We observe that many firms did not grow smoothly, but rather grew (or shrank) rapidly in short, intense phases with jumps in growth rates. 38.5 percent of the observations refer to firms with foreign production. The median share of exports to the EU in sales from domestic production amounts to 35 percent, the corresponding share to countries outside the EU is 15 percent.

Estimation method: The distribution of growth rates does not correspond closely to a normal distribution because of fat long tails; i.e. there are a few firms, growing or shrinking at a very fast pace. So the panel includes a considerable number of extreme values, which would have to be classified as outliers with OLS estimators. Thus, this study takes a robust approach and uses the LAD-estimator or median regressions. The LAD-estimator minimizes the sum of absolute errors and does not give extreme values as much weight as the OLS-estimator. It achieves almost the same efficiency as OLS in situations with independent, but non-normal errors (Hamilton, 1998; see Fiaggio - Konnigs, 1999 for a similar approach). In order to account for possible heteroscedasticity, which would lead to underestimated standard errors by the analytical estimates (Rogers, 1992), bootstrap resampling with 10000 replications is used.

3 The estimation results

The estimation results are provided in four specifications in Table 2. Since the interaction terms induce severe multicollinearity and do not permit the estimation of the full specification (3), four restricted versions were estimated separately.

As found in many studies on the growth of primarily larger firms (see Sutton, 1997, Geroski, 1998 for an overview), there is significant regression to the mean, and Gibrat's law is rejected, indicating that larger firms on average are growing slower in terms of employment. However, the implied speed of adjustment lies between 0.009 and 0.011. So, in line with many other investigations, one can hardly call this (conditional) convergence in size - the implied half-lives vary between 62 and 76 years in specifications I-IV⁶. Rather, during this short time period growth seems to be mainly random and the size differences are persistent.

It is important to control for growth in domestic sales, since - as mentioned above - employment growth originating from domestic sales growth is unlikely to be affected by foreign activities. The significant coefficient of GSHOME indicates that a 1-percent increase induces employment growth of approximately

⁶The speed of convergence is calculated as $b = \frac{1}{T} [\ln(1 + \tau T)]$, where τ is derived from the regressions in Table 2. Half-lives are given by $\ln 2 = b$:

0.41-0.46 percent. The impact of aggregate sales growth is of the same order and confirms the hypothesis that firms operating in growing markets on average exhibit better growth performance. In the present context, the indicator of market growth is rather crude and should be interpreted with care. Firm heterogeneity measured by the average investment sales ratio is also a significant determinant in the all specifications reported in Table 2. Therefore, higher average investment propensity goes hand in hand with faster growth.

Table 2

From a dynamic point of view, the estimates do not suggest significant substitution of domestic production (more precisely, employment embodied in exports) by foreign affiliates, but rather, complementary growth of employment at home and abroad. First, the foreign production dummy is significantly positive (at least at 10%) in Specifications I and II of Table 2; the marginal effect of the foreign production dummy lies between 1.3 and 1.9 percentage points. Secondly, firms which export a large share of their production both to EU-members and to countries outside the EU grow faster on average, according to the estimates of both Specifications II and III. Thirdly, firms pursuing a strategy based on both exports and foreign production grow faster, as shown by the significant interaction effects with foreign production (Specification III). If firms with foreign production exhibit low export orientation, the effect of foreign production is insignificantly negative, but turns positive if exports are sufficiently high. However, estimating interaction terms is - as is often the case - affected by multicollinearity, and thus the parameters could not be estimated precisely. Only the interaction term with exports to countries outside the EU is significant at 10%. The fourth result refers to the persistence of firm size (Specification IV): interacting the foreign production dummy with the initial size of firms results in a significantly positive effect; that is, firm size at the Austrian location is more persistent when the firms operate affiliates abroad and the negative effect of firm size on growth is lower.

The positive relationship between domestic and foreign affiliate employment is also affected by the remaining explaining variables, if they differ systematically between multinational and non-multinational, exporting firms. In particular, a firm's initial size is a candidate here. For this reason, Table 3, compares the median predicted employment growth for several classes of export propensity and calculates the corresponding counterfactual. For firms with no foreign affiliates the foreign production dummy is counterfactually set to one; the other way round is taken for firms with foreign production.

Table 3

We define three groups of exporters: Low export orientation is the category with an EU export share below 20% and a non EU-export share below 10%. To be in the class of high export orientation the corresponding export shares must be larger than 60% and 30%, respectively. The rest is allocated to the middle group.

The evidence seems quite clear: Although firms with foreign production are larger on average ⁷, median predicted growth rates are higher by 0.4-0.5 percentage points. Furthermore, when going through the cells from 'low' to 'high' export orientation, the median predicted employment growth rate increases, especially from the 'low' to 'medium' class of exporters and less from the 'medium' to 'high' grouping (in Specification and II it actually falls slightly). The counterfactuals show a clear, positive impact of running foreign affiliates on domestic employment growth. If firms with no foreign production are classified counterfactually as MNEs, the median predicted growth rate is higher; it is lower if MNEs are classified as pure exporters. With respect to the relationship between foreign and domestic production, it seems decisive that firms achieve high export orientation, especially to countries outside the EU, and that they serve foreign markets according to a strategy, which is built on both high export orientation and foreign production. The substitution of domestic activities by foreign production only seems relevant when firms concentrate solely on shifting production to more favorable foreign locations, stop or reduce exporting, and as a consequence achieve only average or (lower) rates of growth at home. However, even in this case, the counterfactual predictions show that growth rates would be even lower if firms did not produce abroad.

There is not much empirical evidence available for comparison. Buckley et al. (1984) analyze the world's largest firms and find a positive relationship between firm growth and the degree of multinationality (however none for non-US firms). Siddharta - Lall (1982) and Kumar (1984) estimate negative effects for US and British firms. In a more recent paper, Cantwell - Sanna-Randaccio (1992) report a partly contradictory result for the world largest firms⁸. Firms which operate mainly domestically grow faster on average than MNEs because - as the authors argue - of their earlier stage in the internationalization process. In contrast, the degree of multinationality (which could not be tested here) exerts a positive impact on growth performance. However, they do not control for export propensity and contribute to the issue discussed here. Looking at the relationship between export orientation and ex-post growth, the available empirical evidence suggests that fast growing firms are more likely to become exporters. Ex-post, the benefits of exporting are less clear: According to Bernard - Jensen (1999), the employment growth of exporting firms is higher in US manufacturing. However, Wagner - Bernard (1997) and Lui et al. (1999) did not find significant ex-post performance differences between exporters and non-exporters. Concerning the evidence on the common impact of export propensity and foreign production, no evidence so far seems to be available.

4 Conclusions

Based on a cross-section of medium sized to large exporting Austrian manufacturing firms, which participated in the WIFO-Investment Survey, this paper

⁷The Kruskal-Wallis test rejects the hypothesis of equal firm sizes: $\chi^2(1) = 14.85$; $p = 0.0001$.

⁸Note they are by far larger than the firms in the present sample and not really comparable.

investigates in which way export orientation and foreign production are related to the average growth of firms at their domestic location.

Robust median regressions confirm significant regression to the mean: the larger firms grow more slowly on average in terms of employment. However, for firms with production abroad, the negative size effect is smaller. Generally, the effect of firm size on growth is small and one can hardly call this (conditional) convergence in size. Rather, growth seems mainly random and the size differences are persistent

From a dynamic point of view, there seems to be no indication of a substitutional relationship between foreign production and domestic employment growth at the firm level, once it is controlled for domestic sales growth in the regressions. On the contrary, the estimates, as well as the calculation of the counterfactuals, point to positive domestic growth effects of foreign production. Despite their above average size, MNEs based in Austria maintain and sometimes even expand their growth potential at home. In particular, firms which followed a strategy of both high exports (especially with respect to non EU-countries) and foreign production seem to preserve their high growth potential at the domestic location in Austria.

5 References

- Brainard, L.S., A Simple Theory of Multinational Corporations and Trade with a Trade-off between Proximity and Concentration, NBER Working Paper No. 4269, 1993.
- Bernard, A.B., Wagner, J., Exports and Success in German Manufacturing, *Weltwirtschaftliches Archiv* 133(1), 1997, pp. 134-57.
- Bernard, A.B., Jensen, B.J., Exceptional Exporter Performance: Cause, Effect or Both? *Journal of International Economics* 47(1), 1999, pp. 1-25.
- Buckley, B., Dunning, J.H., Pearce, R.D., An Analysis of Growth and Profitability of the World Largest Firms 1972 to 1977, *Kyklos*, 37, 1984, pp. 3-26.
- Buckley, P.J., Casson, M., The Optimal Timing of Foreign Direct Investment, *Economic Journal* 91, 1981, pp. 75-87.
- Cantwell, J.A., Sanna-Randaccio, F., Multinationality and Firm Growth, *Weltwirtschaftliches Archiv* 129(2), 1993, pp. 275-299.
- Dell'mour, R., Egger, P., Gugler, K., Pfaffermayr, M., Wolfmayr-Schnitzer, Y., Outsourcing of Austrian Manufacturing to Eastern Countries: Effects on Productivity and the Labor Market, mimeo, 2000.
- Doms, T., Dunne, T., Roberts, M.J., The Role of Technology Use in the Survival and Growth of Plants, *International Journal of Industrial Organization* 13 (4), 1995, pp. 523-542.
- European Commission, "The Joint Harmonized EU Programme of Business and Consumer Surveys", *European Economy, Reports and Studies* 6, 1997.
- Evans, D.S., Tests of Alternative Theories of the Firm", *Journal of Political Economy* 95(4), 1987, pp. 657-674.
- Faggio, G., Konings, J., Gross Job Flows and Firm Growth in Transition Countries: Evidence Using Firm Level Data on Five Countries., CPR-Discussion Paper 2261, 1999.
- Geroski, P.A., An Applied Econometrician's View of Large Company Performance, *Review of Industrial Organization* 13(3), 1998, pp. 271-293.
- Geroski, P. A., Machin, St.M, Walters, Ch.F., Corporate Growth and Profitability, *Journal of Industrial Economics* 45(2), June 1997, pp. 171-189.
- Hamilton, L., *Statistics with Stata 5.*, Pacific Grove, CA., Duxbury Press, 1998.

- Helpman, E. Krugman, P., *Market Structure and Foreign Trade*, Cambridge, Mass., 1985.
- Jovanovic, B., Selection and the Evolution of Industry, *Econometrica* 50(3), May 1982, pp. 649-670.
- Kumar, M., *Growth, Acquisition and Investment, An Analysis of the Growth of Industrial Firms and their Overseas Activity*, Cambridge, 1984.
- Lui, J.T., Tsou, M.W, Hammit, J.K, Do small plants grow faster? Evidence from the Taiwan Electronics Industry, *Economics Letters* 65, 1999, pp. 121-129.
- Markusen, J.R., The Boundaries of Multinational Enterprises and the Theory of International Trade, *Journal of Economic Perspectives* 9, 1995, pp.169-189.
- Markusen, J.R., Venables, A.J., Multinational Firms and the new trade theory, *Journal of International Economics* 46(2), 1998, pp. 183-203.
- Marries, R., Wood, A. (eds.), *The Corporate Economy*, Macmillan, London 1971.
- Mitchell, M.F., The Scope and Organization of Production: Firm Dynamics over the Learning Curve, *Rand Journal of Economics* 31(1), 2000, pp. 180-205.
- Pfaffermayr, M., Multinationals, Production Externalities, and Complementarity between Domestic and Foreign Activities?, *Swiss Journal of Economics and Statistics* 133(4), 1997, pp. 473-690.
- Rogers, W.H., Quantile Regression Errors, *Stata Technical Bulletin* 9, 192, pp. 16-19.
- Scherer, F.M., Beckenstein, A., Kaufer, E., Murphy, R.D., *The Economics of Multi-Plant Operation, An International Comparison Study*, Harvard University Press, Cambridge MA., 1975.
- Sidharthan, N.S., Lall, S., Recent Growth of U.S. Multinationals, *Oxford Bulletin of Economics and Statistics* 44, 1982, pp. 1-13.
- Sutton, J., "Gibrat's Legacy", *Journal of Economic Literature* 35(1), 1997, pp. 40-59.
- Wagner, J., Bernard, A.B., Exports and Success in German Manufacturing, *Weltwirtschaftliches Archiv*, 133(1), 1997, pp. 134-57.

Table 1: Variable definitions and summary statistics 1996-1999

Variable	Definition	Median	1.Quart.	3.Quart.
$Dln(B)$	Average growth in employment 1996-1999, 100*log differences	0.26	-3.63	3.25
B	Employment 1996	309.30	173.00	581.25
$EXEU^{a)}$	Exports to the EU, as a share of sales from domestic production in 1996	0.35	0.15	0.65
$EXNEU^{a)}$	Exports to countries outside the EU, as a share of sales from domestic production in 1996	0.15	0.05	0.25
$ln(INQ)$	Average investment, as share of sales 1996-1999, both for the home location	5.08	2.71	7.61
$Dln(SHOME)$	Average growth of sales from domestic production*(1-EXEU-EXNEU), 1996-1999	0.55	-0.42	2.48
$Dln(SAGG)$	Average aggregate growth of sales from domestic production (median of the 2-digit industry) using the entire sample of 1144 firms, 1996-1999	3.34		
		Share	Number	
		in%	of firms	
$F^{a)}$	1 if foreign affiliate sales in % of sales from domestic production >0 in 1996, otherwise 0	38.5	135	
$I1^{b)}$	Food, tobacco (15,16), base category	8.6	30	
$I2$	Textiles (17)	5.1	18	
$I3$	Apparel (18)	1.4	5	
$I4$	Leather (19)	1.7	6	
$I5$	Wood, products of wood (20)	2.9	10	
$I6$	Pulp, paper, paper products (21)	6.8	24	
$I7$	Publishing, printing and reproduction	1.7	6	
$I8$	Petroleum, chemicals, (23,24)	8.8	31	
$I9$	Rubber, plastics (25)	5.4	19	
$I10$	Non-metallic mineral products (26)	6.8	24	
$I11$	Basic (27)	8.6	30	
$I12$	Fabricated metal products (28)	10.3	36	
$I13$	Machinery and equipment (29)	11.1	39	
$I14$	Electrical machinery (31)	6.0	21	
$I15$	Radio, TV and communication equip. (32)	2.8	10	
$I16$	Medical and precision instruments, office machinery (31, 30)	1.4	5	
$I17$	Motor vehicles, other transport equipment (34, 35)	4.8	17	
$I18$	Furniture, manufacturing, n.e.c (36)	5.7	20.00	

Note: Firms with missing values in one of the variables, non-exporters, firms with less than 100 employees and 11 observations for firms with extreme growth rates have been excluded. The sample includes firms in NACE14-NACE36 (in Nace 37 there is only 1 firm which is skipped). The total sample the comprises 351 firms. All variables are from the WIFO-Investment Survey (in collaboration with the EU).

a) For a few firms values for 1997 have been imputed

b) A few NACE-2 digits are merged with others because they included only 1 firm.

Table 2: Robust regression (least absolute deviations) for firm growth at the domestic location in terms of employment

Dependent variable: $\emptyset D \ln(B)$	I		II		III		IV	
	β	$t^{a)}$	β	$t^{a)}$	β	$t^{b)}$	β	$t^{a)}$
$\ln(B_{1996})$	-0.009	-2.3 **	-0.011	-2.9 **	-0.010	-2.6 **	-0.010	-2.6 **
<i>Implied Half-life</i>	76.0		62.0		68.3		68.3	
$\ln(B_{1996}) * F_{1996}$	-	-	-	-	-	-	0.002	2.3 **
F_{1996}	0.013	2.0 **	0.019	1.8 *	-0.009	-0.7	-	-
$EXEU_{1996}$	-	-	0.028	2.0 **	-	-	-	-
$EXNEU_{1996}$	-	-	0.038	1.9 *	-	-	-	-
$EXEU_{1996} * F_{1996}$	-	-	-	-	0.026	1.2	-	-
$EXNEU_{1996} * F_{1996}$	-	-	-	-	0.048	1.7 *	-	-
$\emptyset D \ln(SHOME)$	0.460	4.9 **	0.509	5.2 **	0.456	6.9 **	0.451	4.8 **
$\emptyset D \ln(SAGG)$	0.431	3.2 **	0.340	2.5 **	0.410	3.1 **	0.443	3.3 **
$\emptyset \ln(INQ)$	0.013	2.6 **	0.015	3.1 **	0.016	3.6 **	0.015	2.9 **
N	351							
Goodness of fit- Pseudo $R^{2c)}$	0.17		0.18		0.17		0.17	
F-tests								
Regional-dummies (8, 321)	2.32 p=0.02		2.51 p=0.01		1.86 p=0.07		2.38 p=0.02	
$EXEU_{1996}, EXNEU_{1996}=0$ (2, 416)	-		3.46 p=0.03		-		-	
$EXEU_{1996} * F_{1996}, EXNEU_{1996} * F_{1996}=0$ (2, 416)	-		-		1.97 p=0.14		-	

Note regional dummies (8) and the constant are not reported.

a) Based on the estimated variance covariance matrix obtained by bootstrapping with 10000 replications. Note analytic estimates tend to understate the standard errors in case of heteroscedasticity (Rogers, 1992).

b) Based on analytic estimates of the standard errors as bootstrapping did not yield robust results.

c) Pseudo- R^2 calculated as $1 - (\text{sum of weighted deviations from estimated median}) / (\text{sum of weighted deviations from raw estimate})$

* significant at 10%

** significant at 5%

Table 3: Median of predicted growth rates in percent by type of firm and the counterfactual

Export-Intensity	I		II		III		IV	
	Pre-diction	Counter-factual	Pre-diction	Counter-factual	Pre-diction	Counter-factual	Pre-diction	Counter-factual
<i>no foreign production</i>								
low	-1.1	0.3	-2.1	-0.9	-1.1	-1.3	-1.2	0.2
medium	0.1	1.4	-0.3	0.9	0.3	0.8	0.0	1.4
high	0.0	1.3	0.7	1.9	0.0	1.9	-0.1	1.3
Total	-0.1	1.3	0.0	1.2	-0.1	1.3	-0.1	1.2
<i>foreign production</i>								
low	-0.1	-1.4	-1.3	-2.5	-1.9	-1.6	0.2	-1.8
medium	0.5	-0.8	0.3	-0.9	0.3	-0.9	0.6	-1.0
high	0.3	-1.5	0.7	-0.5	0.8	-1.1	0.3	-1.2
Total	0.3	-0.1	0.4	-0.8	0.3	-1.1	0.4	-1.1

Note: Exports are measured as the share in sales from domestic production. Exports are 'low' if the share of exports to the EU is lower than 20%, and that to non-EU countries is lower than 10%. Firms are allocated to the class of high exports, if the corresponding average-shares are above 60% and 30%, respectively. The rest is in the medium category. Each cell exhibits the median predicted values. The counterfactuals are the median predictions, which result from changing the foreign production dummy.